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Mariano Rabassa

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Nutrición y Rendimiento Escolar: ¿Es Posible Mostrar una Relación Causal Usando Datos de Corte Transversal?

DIRECTOR

Leonardo Gasparini

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Nutrition and education in childhood: is it possible to show a causal relationship using cross-section data?

Mariano Javier Rabassa

Tesis de Maestría Maestría en Economía Universidad Nacional de La Plata

Director de Tesis: Dr. Leonardo Gasparini

Abstract

The fact that nutrition affects education outcomes is accepted by researchers and by policy makers. It is simple. Children cannot learn if they are hungry. The validity of the empirical approaches used to show a causal relationship from nutrition to education is an issue of debate. The presence of unobserved characteristics that influence both variables is the main concern of researchers. The goal of this paper is to study the possibility of overcoming these difficulties using the NHANES III (1988-1994), a cross-section data set with national representation in the US. A set of school outcomes and a dummy that accounts for the "food-insecurity" condition of each child's family are the central variables here. Based on a IVs procedure, it looks for variables that can be used as instruments for the "food-insecurity" condition. The preliminary results indicate that child's height and mother's body mass index are no good instruments to do so. Further research in needed to construct other variables that might turn to be good instruments for food-insecurity.

Key words: nutrition, education, instrumental variables, exogeniety

JEL classification: I2

1. Introduction

The fact that nutrition status affects learning has already been accepted by researchers of different fields and by policy makers. It is simple. You cannot learn if you are hungry. Moreover, researchers in the medical science have proved that even if you are not starving, you can learn more (or do better in your tests) if you eat just before you learn (take a test), and also, that there are some kind of nutrients that facilitate learning more than others.

The existence of this link and its quantification is relevant in terms of public policies, and is particularly important in developing countries that face a very limited budget to finance social programs as education and nutrition. This topic is also relevant for developed countries as the US, where there still are poor children, and where the quality of the food children eat is an increasing concern of policy makers.

Economists from different fields are recognizing the importance of these interactions. Development economists consider that this is one of the channels through which health affects income, and therefore poverty status. Labor economists are concerned about the impact that health, through education outcomes, has on life time earnings even in developed countries such as the US. Public economists are using this link, for example, to show that schools take advantage of it, by changing the nutrition intake of children on test days, to increase the school pass proportion.

In practice, the link between nutrition and cognition is accepted. In the United States there are several programs designed to improve food intakes of school children. Some programs are oriented toward low income students, but others are broader and just address the fact that children learn more when they are not hungry. Also, given the increasing proportion of obese children in US, these programs can also be considered as means to improve the quality of the food children eat.

The biggest program in the United States is the School Breakfast Program (SBP), conducted by the US Department of Agriculture since 1966.¹ The same Department also runs the National School Lunch Program, which operates in a similar way to the SBP.

In addition, almost every state has its own nutritional program, implemented through the school system, and they all stress the fact that food is essential for learning. For example, the Maryland Meals for Achievement (MMFA) is a classroom breakfast project that started in Fall 1998 in several Maryland elementary schools. Participating schools offer free school breakfast in the classroom each morning, regardless of family income.

There is a vast literature that addresses the correlation between nutrition status and education. The definitions used to characterize the relevant variables as well as the undertaken methodologies vary across these studies, depending upon the research question asked and the available data.

The differences in research questions can be viewed as timing differences: short versus long-run effects. The short-run approach concentrates in the effect of food intakes today on some measure of school outcome today, such as math or reading test scores. The long-run approach considers the impact of the accumulation of a poor nutritional status during school years and the final school outcome. In the middle, one may also study the consequences of poor nutrition accumulated at a particular school age (e.g. from birth to the age of 10) on the school outcome at that moment (a test score achieved at age 10).

In this sense, the precise definition of nutrition and education not only depends on the available data but also on the timing approach. Nutrition proxies may include different health conditions, such as anemia or body mass index measures, or more direct questions as "do you have enough food to eat?" There are several proxies of education outcomes. Academic outcomes include math and reading test scores, repetition grade, days absent from school; cognitive outcomes concentrate on performance and aptitude tests; and

¹ The School Breakfast Program is a federally assisted meal program operating in public and nonprofit private schools and residential child care institutions. It began as a pilot project in 1966, and was made permanent in 1975. The School Breakfast Program is administered at the Federal level by the Food and Nutrition Service. At the State level, the program is usually administered by State education agencies, which operate the program through agreements with local school food authorities in more than 78,000 schools and institutions. For a review of this program, see Kennedy and Davis (1998).

psychosocial outcomes include variables such as ever suspended from school, number of good friends, shyness, among others.

This paper uses the 3rd National Health and Nutrition Examination Survey (NHANES III) which is a cross section survey that contains information about several characteristics of health condition of children and their parents, as well as a variety of education outcomes (as described above) and the direct question mentioned before: do you have enough food to eat? In this sense, the analysis presented here may be view as an attempt to measure the short-run effect of nutrition on education, controlling for some long-run measures of nutrition status, among many other things.

Previous results are consistent in finding that better nutrition status is associated with better educational outcomes. However, this correlation cannot be directly translated into a causal relationship from nutrition to education. The presence of unobserved characteristics of parents that influence both variables is the obstacle to doing so. Some studies are aware of this problem and try to control for endogeneity and omitted variable bias, but they are still subject to other identification problems.

Therefore, the question is: Is it possible to show a causal relationship from nutrition to education using cross-section data? The goal of this paper is to analyze this possibility; the key point here is to select some plausible instruments that satisfy the basic conditions to be "good instruments". The survey contains variables that might appear reasonable instruments, but it is necessary to analyze them and to test whether they satisfy the exclusion restriction to avoid reverse causality problems.

The rest of the paper is organized as follows. Section 2 summarizes some background literature in the topic. Section 3 describes the theoretical model and the empirical implementation weaknesses that tries to overcome. Section 4 presents the data and descriptive statistics, and specifies the role that each variable plays in the model estimation. Section 5 shows what happens when "bad instruments" are used in the estimation. Finally, the last section concludes and includes a discussion of other potentially plausible instrument that might be constructed using the NHANES and that will be subject of future research.

2. Background literature

Nutritionists agree upon a causal effect from nutrition to cognition and there is a huge literature using several data sets and variables to prove that. Katherine Alaimo et al. (2001) and Lou Hicks et al. (1982) are two examples of many that could be found. There are also an increasing number of economic studies that analyze this issue. Only three of them are summarized here.

Katherine Alaimo et al. (2001) use the 3rd National Health and Nutrition Examination Survey (NHANES III) to analyze the relationship between food insufficiency and school outcomes. They define a binary variable for food insufficiency based on the answers to the question "Do you have enough food to eat?" and use information about academic (math and reading scores, grade repeated, absenteeism), cognitive (perceptual organization and free-from-distractibility examinations) and psychosocial (ever seen a psychologist, ever suspended, how gets along with other children, among others) outcomes.

They divide a sample of 5,349 children into two age groups of 6 to 11 and 12 to 16 years and estimate each school outcome as a function of the food insufficiency dummy and a set of family and individual variables, including sex, age, metropolitan region, poverty status, race-ethnicity, family head education, employment and marital status, health insurance, family size, previous and current health status, among others. They find that food insufficiency is significatively negatively related to school outcomes. The weakness of this result relies on its empirical approach: they improve upon the simple OLS regression by controlling for variables that may be proxies of the unobserved parents' characteristics that might be influencing both nutrition and school outcomes, but they do not take into account the endogeneity problem explicitly.²

Lou Hicks et al. (1982) examine the possible impact of a supplemental food program provided during the perinatal period upon cognitive functioning approximately five to

 $^{^{2}}$ This problem is similar to the one faced in the estimation of returns to education: there might be some benefit in terms of bias from including family background variables in the regression of earning on schooling, but this certainly is not a bias-free result. See Card (1999) for a review of these issues.

seven years later.³ They use a natural experiment in three rural Louisiana counties where the national Special Supplemental Food Program for Women, Infants, and Children (WIC) was implemented in early 1974. They follow a siblings approach where both siblings received WIC benefits, but one of them entered into the program after one year of age (late supplement – control group) while the other received it during the third trimester of pregnancy, continuing for at least the first year of life (early supplement – treatment group).

The sample has only 27 siblings and is selected based on a "nutritional risk criteria" that includes variables such as a history of anemia or low weight for height in the case of children, or frequent pregnancy, young or old maternal age, anemia, or high weight for height in the case of pregnant women. The educational variables include academic (school grades (in arithmetic, writing and reading) and cognitive (similar to those used by Alaimo et al.) outcomes. They collect information on health measures (height, weight, hemoglobin) from public health clinic records. The results indicate that there is a significant enhancement of most intellectual measures. They discuss the possibility that these results may reflect group differences on unobserved variables rather than nutrition: given that all these children come from low income families (i.e. they meet the criteria for WIC benefits), they suggest that older siblings are more exposed to an environment that might be deprived socially, economically and in other ability-relevant dimensions.

This is an interesting approach in two ways, it makes use of a policy change (the beginning of the WIC program) as a natural way to produce treatment and control groups; and it uses siblings as a control group with the advantage of controlling for family fixed effects. Unfortunately, these natural experiments cannot be produced by independent researchers. The alternative approach is to design and implement random experiments. This is particularly possible in developing countries, where there is much more room to implement new programs and where the inexpensive cost of living makes these program implementations affordable. Taking this into account, some researchers are using the presence of Non-Governmental Organizations (NGOs) to conduct longitudinal

 $^{^{3}}$ The portion of the perinatal period during which special supplementation occurred corresponded to the brain "growth spurt" period, that occurs during the final trimester of pregnancy and a postnatal period of 6 to 12 months.

experiments in poor countries like Kenya or India to analyze different topics as the one investigated here.

In this line of study, Edward Miguel and Michael Kramer (2003) examine the impact of a program in which seventy-five rural Kenyan primary schools where face into deworming treatment in a randomized order during 1997 – 1999. The randomization is applied to schools rather than to children and this constitutes the main difference of their approach from others and they have a strong argument in favor of it. The deworming treatment generates positive externalities from reduced disease transmission. Therefore, if randomization is based on individuals within schools, the control group may benefit from the deworming of treated classmates. By randomizing across schools, they get rid of this threat. They find evidence that the program reduced school absenteeism in treatment schools by one-quarter but they don't find any evidence of a positive effect on test scores.

The impact of health on education is also relevant in terms of future earnings. There are articles that consider the relationship between poor health and low income in adulthood, but these studies face an identification problem: which cause the other? By focusing on children Case, Lubotsky and Paxton (2002) eliminate the channel that runs from health to income. Their study relates lower income families with poor health in childhood, addressing that the effects of poor health accumulate over children's lives, generating an intergenerational transmission of income status from parents to their children. They use several large, nationally representative data sets -among which it is the NHANES III-, and concentrate the analysis on the arrival and impact of chronic health conditions in childhood. Their finding indicate that children's health is positively related to household income, and that the relationship between household income and children's health status becomes more pronounced as children grow older.

David Figlio and Joshua Winicki (2002) use a random sample of 23 Virginia school district in 2000 to test two hypotheses: (i) schools threatened with sanctions due to low performance of their students on standardized tests, "game" the accountability system by altering the nutrient content of school lunches around test days to improve the performance of their students; and (ii) this manipulation of school nutrition on test days is successful to increase test scores.

Nutritionists have proven the existence of "empty calories" that have substantial veryshort-term cognitive effects but no long-term effects. Figlio and Winicki have information about the composition of school lunches in terms of calories, protein, calcium and vitamins A and C on a daily basis. Hey also have information about test days and scores, and they know which schools have problems in terms of their passing rate. The authors find evidence that schools threatened by the accountability system respond by providing students with "empty calories" on test days. Also, for fifth grade students, they find a positive and significatively effect of calorie differential on test days on pass rates.

Although they are not controlling for other means of "gaming" the system, like shaping the test pool (reporting as disable children that are not), or having teachers to cheat, the "empty calorie experiment" is conducted at school over a very short duration, and with no particular warning. Also, it is clearly exogenous with respect to the child and his observed and unobserved characteristics. They also use district fixed effects, which should help control for, e.g., how aggressive the district is on actively influencing test scores.

3. Theoretical framework

This model is based on the pioneer work of Behrman, Rosenzweig, and Taubman (1994), and to some extent is similar to the one used by Miguel and Kramer (2003).

The school outcome of child *i* in family *j* (S_{ij}) is modeled as a function of "food insecurity condition" (H_{ij}) a set of exogenous variables X_{ij} that includes variables characterizing family background (FB_j), which are invariant within family *j*, and other individual characteristics (I_{ij}), and his own "luck", v_{ij} .

$$\ln S_{ij} = \alpha_0 + \alpha_1 H_{ij} + \beta' X_j + v_{ij} \qquad (1)$$

It would be very valuable to include information about school quality (class size, quality of teachers, etc.), which certainly has some influence on the child learning process. Unfortunately, this information is not available in this data set.

The problem here is that food insecurity is not exogenous. In particular, it is correlated with family background and individual characteristics. The main point of this study relies on identify an exclusion restriction, a variable correlated with food insecurity but uncorrelated with school outcome. This variable (or set of variables) are defined as Z_{ij} .

$$\mathbf{H}_{ij} = \delta_0 + \theta' \mathbf{X}_{ij} + \rho' \mathbf{Z}_{ij} + \varepsilon_{ij} \qquad (2)$$

This specification is trying to incorporate the fact that there might be some family characteristics that are both influencing food insecurity and school outcome. The empirical approach would be an instrumental variable approach, where equation (2) is estimated in the first step and then the estimated food insecurity variable, \hat{H}_{ij} , is used to estimate equation (1).

Note here that H_{ij} is a dummy variable that takes a value of 1 if family *j* does not have enough food to eat, and 0 otherwise. This specification is known as a Linear Probability Model (LPM).

The LPM has some problems when estimated by Ordinary Least Squares (OLS): (i) the error term is not normally distributed, (ii) estimated probabilities are not necessarily bounded between zero and one, and (iii) unbiased OLS estimation implies that heteroscedasticity is present.⁴ There are studies that overcome this situation by means of logit or probit specifications, bounding the predicted dependent variable to the interval 0-

 $\begin{array}{lll} Y_i = X_i\beta + u_i & \text{where } Y \text{ is a dummy variable} \\ \text{In OLS regression: } & E(Y) = X\beta \\ \text{In the binary case:} \\ & E(Y) = 1*[\Pr(Y=1)] + 0*[\Pr(Y=0)] = \Pr(Y=1) \\ \text{so we model } \Pr(Y_i=1) = X_i\beta \\ & \text{Problems:} \\ & \text{i) u can take only two values:} \\ & 1 - X_i\hat{\beta} & \text{or } -X_i\hat{\beta}. \\ & \text{So, errors aren't normal.} \\ & \text{ii) Estimation may predict probabilities } < 0 \text{ or } > 1 \\ & \text{iii) } V(Y) = E(Y)*[1 - E(Y)] = X_i\beta*(1 - X_i\beta) & \text{therefore} \\ & Y \text{ is heterocedastic.} \\ & \text{See Heckman and Macurdy (1986) for further detail.} \end{array}$

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⁴ Formally, the LPM specification and its limitations can be seen in this way:

1. Nonetheless, the LPM still may be the correct model or might be justified for practical reasons.⁵

4. The Data

The data comes from the NHANES III, which is a national cross-section survey collected during a long period (1988 – 1994). It collects data on school outcome (math and reading scores, two cognitive test, repetition grade, absenteeism, behavioral problems), on health, nutrition (do you have enough food to eat?) and a set of family and individual characteristics (family income, education, employment and marital status of family head, sex, age, race, region of residence, among others).

This section presents the descriptive statistics of a set of variables considered relevant to estimate the model. Probably the most relevant part of this section is the discussion about potential instruments in sub-section 4.5.

4.1. The dependent variable, S_{ij} :

The school outcome can be approximated by using different set of variables:

a) Cognitive outcomes: block design and digit span scores, available for children of 5 or more years old

To evaluate cognitive level of children, the NHANES III administered a performance exam and verbal component from the Wechsler Intelligence Scale for Children-Revised (WISC-R) test. The Block Design test is a perceptual organization examination in which children are asked to construct designs out of blocks to match a model. The Digit Span test is a freedom-from-distractibility examination in which children are asked to repeat up to eight digits in forward and reverse directions.

⁵ Horrace and Oaxaca (2003) develop a sequential estimation technique, the Sequential Least Squares (SLS), and prove that it produces better results than OLS, logit, or probit. The method consists on sequential elimination of observations with the obvious disadvantage of reducing the sample size in each step.

The scores of both tests were standardized to a common scale of 0 to 20 for each age based on samples obtained by test developers. 6

b) Academic variables: math and reading scores, repeated grade, days absent, available for children of 5 or more years old

Math and reading scores were obtained using parts of the Wide Range Achievement Test-Revised (WRAT-R). As above, these scores were standardized using a o to 20 scale and were derived for each child relative to his/her age group based on test-specific standardization samples created by the test developers.

The full WISC-R and WRAT-R have been used extensively to assess children's IQ and academic skills. From the subsets reported in the NHANES is not possible to infer IQ or academic skills approximations.

The survey ask if the child ever repeated a grade and which grade/s. With this information a dummy that takes a value of 1 if the child ever repeated any grade is constructed.

Also, the number of days absent from school in the previous 12 months to the survey is included, that variable can take values from 0 to 365 (meaning that the child was out of school that year).

c) Psychological outcomes: ever seen a psychologist, ever suspended, number of good friends, shyness, difficulties to get along with other children

A set of binary variables accounting for different aspects of psychosocial outcomes are considered here. The NHANES ask if the child has ever seen a psychologist or has ever suspended from school. These answers are already dummy variables.

There is a categorical variable accounting for the number of friends each child has: 0, 1-2, 3-5, 6 or more. The dummy here accounts for "no friends".

The survey asks how the child reacts when it comes to meeting new children and making friends. The alternative answers are: shy, average, outgoing. So here the variable takes a value of one if the answer is "shy". It also asks if the child has

⁶ See the Documentation Exam file for further detail.

difficulties to get along with others, and the options are: no difficulty, some, or a lot. The dummy here takes a value of one if the child has some or a lot of difficulties.

Table 1 tabulates these variables for the two sub samples used by Alaimo et al., 6-11 and 12-16 years old. With respect to the academic and cognitive tests, there is information about the language used by children to take these tests. The possibilities were English and Spanish and approximately 96% of the children took the exam in English.

4.2. Food insecurity variable, H_{ij}

The more relevant variable here is "food insecurity", which is an endogenous variable in the estimation of school outcome. Therefore, it becomes the dependent variable in the first step of the estimation (equation (2) above).

The Youth questionnaire asks: "Do you have enough food to eat?", with the following alternatives:

- 1) Enough food to eat
- 2) Sometimes not enough food to eat
- 3) Often not enough to eat
- 8) Blank but applicable

A dummy variable is constructed to distinguish families with or without "food insecurity" by considering = 1 if the answer is (2) or (3), 0 if answer is (1) and missing value if answer is (8). In this way, 1,490 families are identified as food-insecure families and they represent 11% of the total. Table 2 presents the distribution for the full sample and the two child's ages sub-samples. Note that this is a *family* variable, which means that, for example, if a family has one child between 6-11 years old and another between 12-16 years old, this family will be counted in each sub-sample.

4.3. Family background, FB_j

The set of family background variables includes some demographic variables, such as region of residence and metropolitan area; other variables that characterize the family: family size, number of family moves (an increasing number of family moves may negatively affect the child), family head education, employment and marital status, family income during the previous 12 months to the interview, family income the last month, the poverty income ratio, and health insurance coverage and its source.

The expected relationships are as follows. Family size affects negatively both school outcome and food availability, given a certain income level. An increasing number of family moves is a sign of instability that may affect the children negatively. The years of education of the family head, as well as his/her employment status (dummy that takes value 1 if employed in the previous two weeks to the interview) and his/her marital status (dummy that takes value 1 if married and spouse is present), are expected to positively affect the dependent variables under study.

The information regarding income is one of the major weaknesses of this survey. There is no information about actual income. Only some information about income ranges is available and the rate of non-response is significant. The variable that reports the family income during the previous 12 months to the interview is divided into categories that start with "less than 1,000", increases with interval of a thousand dollars until 20,000 and then intervals of 5,000 until 50,000 and it is truncated at that point. The variable that reports the family income during the last month is divided into categories that start with "less than 100", increases with interval of a hundred dollars until 1,500 and then intervals of 300 until 4,000 and it is truncated at that point. The variables reported here input to each family the average income of its correspondent category. Finally, the poverty income ratio is a ratio between the median income of each category income and a poverty threshold (the smaller the ratio, the poorer the family). This variable may be useful here because this is the variable used to determine the eligibility to receive welfare benefits of programs such as the WIC.

Some information about health insurance coverage and its source is also considered. The alternative sources include Medicare, Medicaid, and others (all kinds of health insurance plans except those that pay only for accidents).

4.4. Individual characteristics, I_{ij}

The individual characteristics set contains variables such as sex, age, weight, and raceethnicity, the number of months that receive WIC benefits, low birth weight, mother's age at child's birth, prenatal smoke exposure, and birth complications (the last four are available only for 6-11 children), variables that may vary within family. Note that here "mother's age at child's birth" is not trying to measure a low income status behavior. In that case, the relevant variable would be "mother's age at first birth", which is trying to capture the particular impact of mother's age during pregnancy and birth of each child (so either very young or very old mothers may put on risk baby's health).

Tables 3 and 4 present means and standard errors for all these variables. They are tabulated for each family category in terms of food insecurity, including a two-tail mean difference test. The results are similar for both age groups.

Academic and cognitive outcomes are significatively different between food-insecure and food-secure families, with the exception of days absent from school for the older group of children. The psychosocial outcomes present some differences. While younger children have significant differences with respect to shyness and difficulties to get along with other children, older ones have significant difference regarding suspension from school and lack of friends.

Regarding the set of family background characteristics, there are no differences in the proportion of children living in a metropolitan area. Half of children between 6-11 years old live in a metropolitan area and almost half (47%) of older children also live in a metropolitan area.⁷ There are also no differences in the number of family's moves for both groups.

^{7 &}quot;Metropolitan area" is a dummy variable that takes value 1 if the child lives in central counties of metro areas of 1 million population or more, OR, fringe counties of metro areas of 1 million population or more.

All the other variables are significatively different between the two groups. In particular, families with food insecurity tend to have larger family size; to have a family head less educated, without employment (in the two weeks before the interview) and not married; to lack of health insurance and to be poorer.

Surprisingly, the number of families covered by Medicare or Medicaid is not significatively different between food-insecure and food-secure families in both age groups.

Finally it is the set of individual characteristics. There are no significant differences in terms of sex, age or weight of children in both groups. With respect to the race-ethnicity variables, the proportion of black children is not significatively different between families with or without food insecurity. But there is a significant difference for Hispanic children: within food-insecure families, 55% (49%) of children between 6 and 11 (12 and 16) years old are Hispanic, against 27% (31%) for food-secure families.

Children in food-insecure families tend to receive the WIC benefits for a longer period than the ones in food-secure families.

For children between 6 and 11 years old, there are no significative differences for the dummies of low birth weight and birth complications. Instead, the proportion of mothers who smoked while pregnant is smaller among food-insecure families, at a significance level of 10%. In this group, mothers were much younger at their child's births: 28 against 41 years old. This result might be more associated with the low-income hypothesis of fertility, which states that poorer people tend to have children at younger stages.

4.5. Are there "good instruments" for food-insecurity?

The key point to prove the causal relationship from nutrition to school outcomes is to overcome econometric difficulties, due to endogeneity (reverse causality) and omitted variable (confounding) bias.

In terms of the results, the power of the model relies on its capability to isolate the effect of nutrition on school outcomes. For example, consider the case in which a child with poor nutrition status also tends to have below-average school performance, i.e., in terms of school attendance. One possible interpretation of this observed pattern is that poor health is the cause of higher absenteeism for this child. But a second plausible explanation is that the child comes from households with unobservable lower socioeconomic status, and that this fact leads both to worse educational outcomes and health outcomes, or similarly, the child has parents with unobservable less interest in both their child's health and education, leading to a correlation between child health and education that is not causal. In many cases it is almost impossible to distinguish between these two explanations, especially when the data used in the analysis consists of a single cross-section of observations, as the NHANES III is. One may ask why to work with this data set. The answer is that the NHANES III is the only source of both health and school outcome information about children at a national level (and public access).

The variable used to construct the dummy of food-insecurity condition does not vary within families. This eliminates the use of siblings to control for family fixed effects.

Many of the family background variables influence food-insecurity as well as school outcomes, such as family income, family head education, employment and marital status, among others.

At the moment, it is not possible to say that there is a good instrument for food insecurity. At first sight, the following two instruments appear as potentially reasonable.

First, child's height. There is a well known positive correlation between nutrition status and individual's height. The data used here supports this fact: table 4 reports that children in food-secure families are significatively taller than the other ones. Disregarding this fact, there is no reason to expect correlation between height and school outcome. Unfortunately, as table 5 shows, for children between 6-11 years old the only dependent variable not significatively correlated with height is the digit span score; while for the older group reading scores and grade repetition are the ones uncorrelated with height.

Second, the body mass index of parents. This variable might be correlated with food insecurity and uncorrelated with children's school outcomes. From the data available in the NHANES it is possible to construct a BMI index for parents. In fact, this variable is already computed for all surveyed individuals, but it is not possible to find the relationship between children and their parents. Instead, there is information about

mother and father's weight and height and they can be used to compute the BMI in the following way:⁸

 $BMI = \frac{weight (in kg)}{\begin{pmatrix}height (in cm)/100\end{pmatrix}^2}.$

Table 4 presents the statistics for mother and father's BMI. In both age groups the average mother's BMI is higher for families with food insecurity, while the average father's BMI index is not significatively different between the two groups. One possible explanation of the higher average BMI among mothers with food-insecurity condition is that there might be a higher proportion of obese mothers, which can be considered as a sign of low quality food.⁹

The correlation of mother's BMI and food-insecurity is very strong, as can be seen in Table 5. Regrettable, it is also correlated, at a 10% significance level with the whole set of dependent variables. On the contrary, father's BMI appears to have no relation not only with school outcome variables but also with the food-insecurity one.

These results suggest that none of them can be considered as "good" instruments for food insecurity. Next section presents some regressions using these "bad instruments". Although the results appear to be good, it is not possible to draw conclusions from it, as the ones presented in Alaimo et al. (2001). Finally, the last section includes other potentially plausible instrument that might be constructed using the NHANES and that will be subject of future research.

⁸ This formula is the one used by NHANES to compute the variable BMI and it is based in centimeters and kilograms. The variables for mother and father's height and weight are expressed in inches and pounds, so will convert them using the following scale: 1 inch = 2.54 cm, 1 pound = .4536 kg

⁹ Studies discussing the relation of obesity and poverty in the US are receiving growing attention among economic researchers.

5. Empirical results using "bad" instruments

At this stage of the analysis, it seems that the NHANES III does not contain information to find empirical evidence of a causal relationship between food-insecurity and school outcomes. For the sake of simplicity, this section only includes results for one dependent variable, "math scaled scores". Although there are variations using different dependent variables, the point here is to address the quality of the instruments, which is similar with respect to the reminding dependent variables.

First, it might be interesting to see the simple OLS estimation of school outcomes on food-insecurity. Table 6 reports a significant relationship between these variables; even controlling for some family and individual characteristics, this relationship remains significatively different from zero. But it has been argued here that these estimations cannot be interpreted as a causal relationship because of the presence of unobservables that affect both food-insecurity and school outcome. The problem does not end up there. It is necessary to find a good instrument to eliminate this endogeneity problem.

The first step is to estimate equation (2), which has a binary variable as dependent variable. Table 7 presents the results based on a linear probability model for food-insecurity. Two of the instruments presented in section 4.5 (child's height and mother's BMI) are used alternatively as explanatory variables for the two age sub-samples. The results indicate that these variables are not good enough in explaining families' food-insecurity. Only in model (6), where mother's BMI is used to instrument food-insecurity among children of 6 to 11 years old, this variable turns to be significatively different from zero. But as seen in the correlation table, it is not exogenous with respect to school outcome.

However, when using these estimates in the second step (equation (1) of the theoretical model), Table 8 shows that there is a strong and negatively significant relationship between food insecurity and school outcomes (in particular, math scores).

This is the source of criticism to many nutritional studies as the one of Alaimo et al (2001). In spite of the efforts made in order to control for unobserved characteristics that affect both nutrition and education variables -generating endogeneity and omitted

variable bias-, the present study shows that with the tools provided by NHANES III it is not possible to identify a good instrument to account for these empirical problems.

6. Conclusions and future research

Is it possible to show a causal relationship between nutrition and education using crosssection data? At the moment, the answer is "NO".

Although the empirical results can be viewed by a careless reader as evidence in favor of this relationship, they are still subject to endogeneity and omitted variable bias and therefore inferences cannot be made. This is not only true for this study, but also for the results of Alaimo et al. (among others) that face the same limitations as the ones detailed here.

At this point, there are two alternatives. One is to consider that these results are enough evidence to conclude that this empirical strategy is inappropriate and should be avoided. The other is to keep on trying to check if there are any other variables in the survey that can be used to overcome these identification problems. Given that there are no many other ways to study this link, at least, using public data sets with national representation, it might be worth while to insist in this path of work.

In fact, there are already identified other instruments which can be developed.

One possibility would be to construct an exogenous state policy variable that relates the availability of different levels of food assistance at school with the use of these programs by children from food-insecure families. The construction of this variable would take some steps.

First, the NHANES III contains some information about state of residence: there is a variable that indicate the state of residence if the family lives in a county with a population of more than 500,000 habitants. Therefore, one might narrow the analysis to children living in metropolitan areas. Second, there is also information about breakfast and lunch served at school at a fixed price, and also the frequency (days per week) that the children have breakfast or lunch at school. Table 9 presents the description of these

four variables –and the results show that there is a significative difference in the use of these programs between food-secure and food-insecure children. Third, the survey was conducted between 1988 and 1994. Although there is no information about the exact moment at which each interview took place, there are some phases and sub-phases of data collection. Fourth, it would be necessary to collect from other sources the existence of state school breakfast and lunch programs, their coverage and their year of implementation (and further extensions). Combining all these variables it could be possible to find variation between states and within states for different periods.

Another possibility is the use of the poverty-income ratio to determine the eligibility of families to programs such as the WIC. Together with other eligibility criteria, this information could be use to construct a variable that accounts for the availability and effective use of this kind of public programs.

If after these and other attempts is still unattainable to find a good instrument for foodinsecurity, i.e., if this "good instrument" does not exist, the results of this study would imply that it is not possible to use cross-section data sets, in particular, the NHANES III, to show a causal relationship from nutrition to education.

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Table 1
Descriptive statistic of the dependent variable: school outcomes, 1988-1994
I. Children between 6 and 11 years old

I. Children between 6 and 11 years old					
Variable	Obs	Mean	Std. Dev.	Min	Max
Academic outcomes					
Math - scaled scores	3,139	7.89	3.34	0	20
Reading - scaled scores	3,129	6.85	3.78	0	18
Ever repeated grade (=1 if yes)	3,317	0.16	0.37	0	1
Days absent last year	3,434	4.56	7.99	0	240
Cognitive outcomes					
Block design - scaled score	3,119	8.60	3.34	1	19
Digit span - scaled score	3,116	8.16	3.00	1	19
Language used in exams					
1 if English, 0 if Spanish	3,134	0.95	0.21	0	1
Psychosocial outcomes					
Ever seen a psychologist (=1 if yes)	3,461	0.07	0.26	0	1
Ever suspended from school (=1 if yes)	3,318	0.04	0.21	0	1
Has no friends (=1 if yes)	4,720	0.03	0.17	0	1
Shy (=1 if yes)	3,461	0.24	0.42	0	1
Difficulties to get along (=1 if yes)	3,461	0.02	0.15	0	1
II. Children between 12 and 16 years old					
Variable	Obs	Mean	Std. Dev.	Min	Max
Academic outcomes					
Math - scaled scores	1,940	7.17	3.40	0	20
Reading - scaled scores	1,931	7.39	3.60	0	17
Ever repeated grade (=1 if yes)	2,210	0.28	0.45	0	1
Days absent last year	2,182	8.12	24.81	0	365
Cognitive outcomes					
Block design - scaled score	1,915	8.19	3.41	1	19
Digit span - scaled score	1,915	7.72	2.85	1	19
Language used in exams					
1 if English, 0 if Spanish	1,939	0.97	0.18	0	1
Psychosocial outcomes					
Ever seen a psychologist (=1 if yes)	2,211	0.13	0.34	0	1
Ever suspended (=1 if yes)	2,210	0.21	0.41	0	1
Has no friends (=1 if yes)	2,216	0.03	0.18	0	1
Shy (=1 if yes)	2,213	0.23	0.42	0	1
Difficulties to get along (=1 if yes)	2,212	0.02	0.15	0	1

Note: Cognitive function was evaluated using parts of two tests, the Wechsler Intelligence Scale for Children-Revised (WISC-R) and the Wide Range Achievement Test-Revised (WRAT-R). Two subtests of the WISC-R test, a verbal component (Digit Span) and a performance exam (Block Design), were administered and are considered relatively culturally unbiased (Kaufman, 1979). In addition, two subtests of the WRAT-R test, math and reading, were conducted. The WISC-R test was administered first and was followed by the WRAT-R. The scores for all four subtests used a common scale and were derived for each child relative to his/her age group based on test-specific standardization samples created by the test developers (Wechsler, 1974; Jastak, 1984; Kramer, 1995).

Source: NHANES III Exam Data and Documentation Files

Table 2	
Families and food insecurity: <i>Do you have enough food to eat?</i>	

	I. Full Sample		II. Children 6-11		III. Children 12-16	
	Obs	%	Obs	%	Obs	%
Yes	12,407	89.28	4,207	89.43	1,967	89.04
No	1,490	10.72	497	10.57	242	10.96
Total	13,897	100.00	4,704	100.00	2,209	100.00

Note: the answer "no" includes families that respond "Sometimes not enough food to eat" or "Often not enough food to eat"

Source: NHANES III Youth Data File

Table 3
Descriptive statistic of dependent and family background variables, clasiffied according to "family food insecurity"

¥7 · · · ·	-	en between 6 and 11 year	's old	Children between 12 and 16 years old			
Variable	Family with food insecurity	Family without food insecurity	Pooled	Family with food insecurity	Family without food insecurity	Pooled	
Dependent variable	msceurty	histeunty		msceurry	mseeunty		
Math scaled score	6.76***	8.03***	7.89	5.71***	7.36***	7.17	
	(0.16)	(0.06)	(0.06)	(0.20)	(0.08)	(0.08)	
Reading scaled score	5.36***	7.04***	6.85	5.91***	7.59***	7.40	
	(0.19)	(0.07)	(0.07)	(0.22)	(0.09)	(0.08)	
Ever repeated grade (=1 if yes)	0.26***	0.15***	0.16	0.43***	0.26***	0.28	
	(0.02)	(0.01)	(0.01)	(0.03)	(0.01)	(0.01)	
Days absent last year	5.34**	4.47**	4.57	9.15	8.02	8.14	
	(0.47)	(0.14)	(0.14)	(0.91)	(0.59)	(0.53)	
lock design scaled score	8.05***	8.67***	8.60	7.18***	8.32***	8.19	
	(0.16)	(0.06)	(0.06)	(0.23)	(0.08)	(0.08)	
igit span scaled score	7.27***	8.27***	8.16	6.81***	7.84***	7.72	
	(0.15)	(0.06)	(0.05)	(0.20)	(0.07)	(0.07)	
ver seen a psychologist (=1 if yes)	0.08	0.07	0.07	0.15	0.13	0.13	
	(0.01)	(0.00)	(0.00)	(0.02)	(0.01)	(0.01)	
ver suspended from school (=1 if yes)	0.07	0.04	0.04	0.34***	0.20***	0.22	
	(0.01)	(0.00)	(0.00)	(0.03)	(0.01)	(0.01)	
as no friends (=1 if yes)	0.04	0.03	0.03	0.07***	0.03***	0.03	
	(0.01)	(0.00)	(0.00)	(0.02)	(0.00)	(0.00)	
hy (=1 if yes)	0.28**	0.23**	0.24	0.32***	0.22***	0.23	
	(0.02)	(0.01)	(0.01)	(0.03)	(0.01)	(0.01)	
Difficulties to get along (=1 if yes)	0.04***	0.02***	0.02	0.03	0.02	0.02	
	(0.01)	(0.00)	(0.00)	(0.01)	(0.00)	(0.00)	
amily Background							
ive in Metropolitan Area (=1 if yes)	0.49	0.50	0.50	0.43	0.48	0.47	
	(0.02)	(0.01)	(0.01)	(0.03)	(0.01)	(0.01)	
ive in Northeast region (=1 if yes)	0.085** (0.01)	0.12** (0.00)	0.11 (0.00)	0.12 (0.02)	0.11 (0.01)	0.11 (0.01)	
ive in Midwest region (=1 if yes)	0.14***	0.19***	0.19	0.11***	0.17***	0.17	
	(0.02)	(0.01)	(0.01)	(0.02)	(0.01)	(0.01)	
ive in South region (=1 if yes)	0.32***	0.39***	0.39	0.46	0.42	0.42	
	(0.02)	(0.01)	(0.01)	(0.03)	(0.01)	(0.01)	
ive in West region (=1 if yes)	0.42***	0.24***	0.26	0.30**	0.24**	0.01	
	(0.02)	(0.01)	(0.01)	(0.03)	(0.01)	(0.43)	
amily size	5.61***	4.76***	4.85	5.62***	4.91***	4.99	
	(0.10)	(0.03)	(0.03)	(0.13)	(0.04)	(0.04)	
umber of family moves	2.62	2.96	2.91	3.40	3.48	3.47	
	(0.14)	(0.08)	(0.07)	(0.23)	(0.11)	(0.10)	
amily head years of education	8.21***	11.48***	11.14	8.65***	11.08***	10.82	
	(0.19)	(0.06)	(0.06)	(0.26)	(0.09)	(0.08)	
amily head employed (=1 if yes in past 2 weeks)	0.56***	0.78***	0.75	0.51***	0.78***	0.75	
	(0.02)	(0.01)	(0.01)	(0.03)	(0.01)	(0.01)	
amily head married (=1 if yes w/ spouse present)	0.54*** (0.02)	0.68*** (0.01)	0.66 (0.01)	0.43*** (0.03)	0.64*** (0.01)	0.61 (0.01)	
amily income last 12 months (in US\$) (#)	12,055***	26,093***	24,682	11,064***	26,880***	25,242	
	(374)	(250)	(237)	(500)	(376)	(357)	
amily income last month (in US\$) (#)	878***	1,940***	1,831	847***	1,989***	1,865	
	(27)	(20)	(19)	(36)	(30)	(28)	
overty income ratio (#)	0.75*** (0.03)	1.89*** (0.02)	1.77 (0.02)	0.69*** (0.03)	1.89*** (0.03)	1.77 (0.03)	
overed by health insurance last month (=1 if yes)	0.77*** (0.02)	0.87*** (0.01)	0.86 (0.01)	0.67*** (0.03)	0.82*** (0.01)	0.80	
overed by Medicare last month (=1 if yes)	0.13 (0.07)	0.06 (0.02)	0.07 (0.02)	0.13 (0.13)	0.05 (0.02)	0.05 (0.02)	
overed by Medicaid last month (=1 if yes)	0.89 (0.02)	0.86 (0.01)	0.87 (0.01)	0.69 (0.04)	0.76 (0.02)	0.74 (0.02)	
Covered by other health ins. last month (=1 if yes)	0.86** (0.03)	0.92** (0.01)	0.91 (0.01)	0.82** (0.05)	0.90** (0.01)	0.90 (0.01)	
umber of observations (##)	497	4,207	4,704	242	1,967	2,209	

 Autore of observations (##)
 497
 4,207
 4,704
 242
 1,967
 2,209

 Note: Standard errors are in parenthesis. * indicates a statistically significant difference between food-insecure and food-secure families at a 90% confidence interval, #* at a 95% and

 *** at a 99% confidence interval. (#) Income variables are reported in categories. Here it was attached to each variable the mean income of each category. The poverty income and a poverty threshold, i.e., the smaller the ratio, the poorer the family. This variable is subject to nonresponse error. See footnote in text. (##) The total number of observations for each variable may vary due to missing cases.

 Source: NHANES III Youth and Exam data files

Table 4	
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Descriptive statistic for individual characteristics and	possible instrument variables, clasiffied according to "family food insecurity"
Descriptive statistic for mulvidual characteristics and	possible instrument variables, clashifted according to hanny food insecurity

	Children	between 6 and 11 y	ears old	Children l	between 12 and 16 y	ears old
Variable	Family with food insecurity	Family without food insecurity	Pooled	Family with food insecurity	Family without food insecurity	Pooled
Individual characteristics						
Sex (=1 if boy)	0.49	0.51	0.51	0.48	0.47	0.47
	(0.02)	(0.01)	(0.01)	(0.03)	(0.01)	(0.01)
Age (in years)	8.55	8.51	8.51	13.98	13.99	13.98
	(0.08)	(0.03)	(0.03)	(0.09)	(0.03)	(0.03)
Black (=1 if yes)	0.28	0.29	0.29	0.38	0.36	0.36
	(0.02)	(0.01)	(0.01)	(0.03)	(0.01)	(0.01)
Hispanic (=1 if yes)	0.55***	0.27***	0.30	0.49***	0.31***	0.33
	(0.02)	(0.01)	(0.01)	(0.03)	(0.01)	(0.01)
Weight (in lbs)	60.48	59.00	59.16	127.63	129.96	129.69
	(1.46)	(0.51)	(0.48)	(2.29)	(0.82)	(0.77)
How many months received WIC benefits	17.29***	17.28***	17.28	27.71	25.68	25.91
	(0.65)	(0.25)	(0.23)	(1.28)	(0.44)	(0.42)
Mother smoked while pregnant (=1 if yes)	0.17* (0.02)	0.21* (0.01)	0.21 (0.01)			
Mother's age at child's birth	28.10** (2.63)	40.96** (1.87)	39.60 (1.70)			
Low birth weight (= 1 if less than 5 1/2 lbs)	0.09 (0.01)	0.07 (0.00)	0.07 (0.00)			
Birth complications (=1 if yes) (2)	0.12 (0.01)	0.11 (0.00)	0.11 (0.00)			
Possible instruments?						
Child's height (in inches)	52.52*	52.97*	52.92	63.31***	64.01***	63.93
	(0.26)	(0.09)	(0.08)	(0.24)	(0.09)	(0.08)
Mother's BMI	27.25***	25.60***	25.76	28.40***	27.17***	27.29
	(0.29)	(0.09)	(0.09)	(0.46)	(0.15)	(0.14)
Father's BMI	26.15	26.36	26.34	26.91	26.93	26.93
	(0.22)	(0.07)	(0.07)	(0.37)	(0.11)	(0.11)
Number of observations (##)	497	4,207	4,704	242	1,967	2,209

 Mode:
 Standard errors are in parenthesis. * indicates a statistically significant difference between food-insecure and food-secure families at a 90% confidence interval, ** at a 95% and *** at a 95% confidence interval

 (##) The total number of observations for each variable may vary due to missing cases

 Source: NHANES III Youth and Exam data files

Table 5 Partial correlation among relevant variables

	Food-insecure family (=1 if yes)	Health insurance last month (=1 if yes)	Child's height (in inches)	Mother's BMI	Father's BMI	Times per week child eats breakfast at school	Times per week child eats breakfast at school	Observation
Children between 6-11 years old								
Dummy endogenous variable								
Food-insecure family (=1 if yes)	1.000	-0.011	-0.009	0.088	-0.030	0.163	0.029	1,498
		(0.672)	(0.716)	(0.001)	(0.253)	(0.000)	(0.268)	
Dependent variables								
Math - scaled scores	-0.051	0.053	0.074	-0.084	-0.007	-0.160	-0.004	1,428
	(0.055)	(0.046)	(0.005)	(0.002)	(0.798)	(0.000)	(0.887)	
Reading - scaled scores	-0.051	0.120	0.267	-0.126	0.005	-0.210	-0.029	1,420
	(0.053)	(0.000)	(0.000)	(0.000)	(0.844)	(0.000)	(0.271)	
Ever repeated grade (=1 if yes)	0.071	-0.044	0.101	0.093	-0.036	0.119	-0.003	1,451
	(0.007)	(0.097)	(0.000)	(0.000)	(0.170)	(0.000)	(0.911)	
Days absent last year	0.003	0.058	-0.052	0.034	0.018	-0.015	0.002	1,497
	(0.899)	(0.025)	(0.047)	(0.192)	(0.478)	(0.552)	(0.930)	
Block design - scaled score	0.014	0.007	-0.070	-0.070	0.013	-0.214	-0.059	1,416
	(0.611)	(0.798)	(0.009)	(0.008)	(0.632)	(0.000)	(0.027)	
Digit span - scaled score	-0.052	0.067	0.027	-0.076	0.004	-0.090	-0.028	1,417
	(0.049)	(0.012)	(0.320)	(0.005)	(0.891)	(0.001)	(0.301)	
Ever seen a psychologist (=1 if yes)	0.037	0.043	0.082	0.017	-0.017	-0.048	-0.001	1,496
	(0.158)	(0.099)	(0.002)	(0.506)	(0.518)	(0.063)	(0.967)	,
Ever suspended from school (=1 if yes)	0.030	0.011	0.108	0.002	-0.018	0.038	0.025	1,451
Ever suspended from senoor (1 fr yes)	(0.260)	(0.669)	(0.000)	(0.951)	(0.494)	(0.146)	(0.347)	1,451
Has no friends (=1 if yes)	-0.027	-0.007	-0.039	0.046	-0.033	0.018	0.033	1,498
has no mends (-1 myes)								1,498
	(0.303)	(0.775)	(0.132)	(0.078)	(0.201)	(0.485)	(0.203)	1 407
Shy (=1 if yes)	0.014	-0.060	-0.006	0.002	0.028	-0.027	0.029	1,497
	(0.586)	(0.021)	(0.806)	(0.925)	(0.274)	(0.290)	(0.269)	
Difficulties to get along (=1 if yes)	0.078	0.031	-0.025	0.038	-0.022	-0.015	0.022	1,498
	(0.003)	(0.234)	(0.332)	(0.143)	(0.396)	(0.559)	(0.395)	
Children between 12-16 years old								
Dummy endogenous variable								
Food-insecure family (=1 if yes)	1.000	-0.062	-0.034	0.024	-0.048	0.148	0.030	950
		(0.057)	(0.300)	(0.453)	(0.139)	(0.000)	(0.356)	
Dependent variables								
Math - scaled scores	-0.095	0.076	0.095	-0.093	0.027	-0.137	-0.111	897
	(0.004)	(0.024)	(0.004)	(0.005)	(0.430)	(0.000)	(0.001)	
Reading - scaled scores	-0.093	0.108	0.051	-0.089	-0.021	-0.188	-0.047	889
5	(0.006)	(0.001)	(0.131)	(0.008)	(0.527)	(0.000)	(0.159)	
Ever repeated grade (=1 if yes)	0.086	0.000	0.043	0.045	0.006	0.134	0.045	950
	(0.008)	(0.998)	(0.191)	(0.164)	(0.864)	(0.000)	(0.171)	
Days absent last year	0.074	0.004	0.071	0.050	0.032	-0.012	-0.034	946
	(0.023)	(0.908)	(0.030)	(0.129)	(0.335)	(0.725)	(0.292)	
Block design - scaled score	-0.086	0.033	0.083	-0.115	0.023	-0.050	-0.068	887
stock design searce score	(0.011)	(0.331)	(0.014)	(0.001)	(0.507)	(0.141)	(0.044)	007
Digit man seeled seere		0.107	0.066	-0.055	-0.060	-0.099		887
Digit span - scaled score	-0.064			-0.055 (0.107)			-0.062	00/
Ever seen a marchalagist (=1 if ves)	(0.061)	(0.002)	(0.053)		(0.078)	(0.004)	(0.067)	948
Ever seen a psychologist (=1 if yes)	0.072	0.027	0.086	0.011	-0.074	-0.041	-0.100	948
Summer and the second	(0.028)	(0.407)	(0.009)	(0.736)	(0.024)	(0.212)	(0.002)	0.50
Ever suspended from school (=1 if yes)	0.102	0.041	0.159	0.053	0.002	0.085	-0.050	950
	(0.002)	(0.206)	(0.000)	(0.107)	(0.958)	(0.009)	(0.123)	
Has no friends (=1 if yes)	0.003	-0.025	0.009	0.035	-0.010	0.056	-0.044	950
	(0.925)	(0.450)	(0.784)	(0.284)	(0.759)	(0.087)	(0.176)	
Shy (=1 if yes)	0.110	-0.090	-0.072	0.030	0.054	-0.034	-0.007	950
	(0.001)	(0.006)	(0.026)	(0.364)	(0.096)	(0.293)	(0.838)	
Difficulties to get along (=1 if yes)	0.050	0.033	-0.073	0.028	0.040	-0.027	-0.046	950
	(0.129)	(0.317)	(0.025)	(0.396)	(0.215)	(0.409)	(0.156)	

Table 6OLS regression on math scores (in logs)

Variables	Children 6-	Children 12-16 years old		
variables	(1)	(2)	(3)	(4)
Not enough food to eat (=1 if yes)	-0.1786	-0.0739	-0.2519	-0.0987
	(0.0264)	(0.0304)	(0.0376)	(0.0377)
Family income (last 12 mo)		3.76E-06		-0.0201
		(8.02E-07)		(0.0069)
Family size		-0.0146		0.0302
		(0.0058)		(0.0038)
Family head years of education		0.0170		0.0692
		(0.0031)		(0.0299)
Family head employed (=1 if yes in past 2 weeks)		-0.0118		0.1247
		(0.0247)		(0.0267)
Family head married (=1 if yes w/ spouse present)		0.0371		0.1219
		(0.0229)		(0.0307)
Covered by health insurance last month (=1 if yes)		-0.0010		0.0476
		(0.0263)		(0.0232)
Live in Midwest region (=1 if yes)		-0.0074		-0.0281
		(0.0336)		(0.0440)
Live in South region (=1 if yes)		-0.0199		0.0304
		(0.0312)		(0.0406)
Live in West region (=1 if yes)		0.0165		0.0471
		(0.0347)		(0.0447)
Live in Metropolitan Area (=1 if yes)		-0.0120		-0.0048
		(0.0190)		(0.0252)
Child's sex		0.0793		0.3852
		(0.0177)		(0.1932)
Child's age		0.0561		-0.0144
-		(0.0604)		(0.0069)
Child's age squared		-0.0024		-0.1964
		(0.0035)		(0.0304)
Black (=1 if yes)		-0.0837		-0.0634
		(0.0241)		(0.0344)
Hispanic (=1 if yes)		-0.0499		. ,
		(0.0279)		
Mother's age at child's birth		0.0005		
		(0.0014)		
Mother's age at child's birth sq		-3.98E-07		
		(1.43E-06)		
Mother smoked while pregnant (=1 if yes)		-0.0572		
		(0.0222)		
Constant	2.0322	1.4485	1.9074	-1.1015
	(0.0089)	(0.2620)	(0.0127)	(1.3437)
Sample size	3,029	2,468	1,889	1,599
R-squared	0.0149	0.1021	0.0232	0.1798

Notes: standard errors are in parenthesis, regional omitted category: northeast

Table 7
First step LPM estimation: food-insecurity dummy

	Children 6-	11 years old	Children 12-16 years old	
Variables	(5)	(6)	(7)	(8)
Child's height (in inches)	-0.0007		-0.0022	
	(0.0012)		(0.0019)	
Mother's BMI		0.0017		-0.0002
		(0.0008)		(0.0012)
Months received WIC benefits	0.0001	0.0001	0.0005	0.0007
	(0.0004)	(0.0003)	(0.0004)	(0.0004)
Family size	0.0103	0.0088	0.0137	0.0124
	(0.0036)	(0.0030)	(0.0042)	(0.0044)
Family income (last 12 mo)	-3.67E-06	-3.77E-06	-4.79E-06	-4.88E-06
	(4.85E-07)	(3.96E-07)	(5.95E-07)	(6.17E-07)
Family head years of education	-0.0124	-0.0095	-0.0013	-0.0005
	(0.0018)	(0.0016)	(0.0023)	(0.0025)
Family head employed (=1 if yes in past 2 weeks)	-0.0380	-0.0267	-0.0455	-0.0385
	(0.0154)	(0.0128)	(0.0192)	(0.0198)
Family head married (=1 if yes w/ spouse present)	-0.0065	0.0161	-0.0279	-0.0182
	(0.0140)	(0.0116)	(0.0173)	(0.0178)
Covered by health insurance last month (=1 if yes)	-0.0036	-0.0098	-0.0254	-0.0164
	(0.0164)	(0.0148)	(0.0195)	(0.0206)
Constant	0.3529	0.2269	0.3775	0.2194
	(0.0684)	(0.0358)	(0.1268)	(0.0530)
Sample size	2,721	3,482	1,741	1,607
R-squared	0.1091	0.0955	0.1104	0.0989

Notes: standard errors are in parenthesis

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Table	8
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2nd step regression on math scores	(in logs) - not enough food to eat instrumented
2nd step regression on math scores	(In logs) - not enough loou to eat instrumented

· · · · · · · · · · · · · · · · · · ·	Children 6-	food to eat instrumented Children 6-11 years old		Children 12-16 years old	
Variables	(9)	(10)	(11)	(12)	
Not enough food to eat (=1 if yes) (#)	-14.1601	-3.4357	-3.6838	-0.4917	
	(3.9999)	(0.9013)	(1.0700)	(1.0427)	
Family income (last 12 mo)	-4.79E-05	-8.93E-06	-1.36E-05	2.31E-06	
	(1.47E-05)	(3.53E-06)	(5.28E-06)	(5.24E-06)	
Family size	0.1310	0.0175	0.0312	-0.0116	
	(0.0417)	(0.0100)	(0.0166)	(0.0150)	
Family head years of education	-0.1580	-0.0161	0.0199	0.0290	
	(0.0498)	(0.0092)	(0.0043)	(0.0044)	
Family head employed (=1 if yes in past 2 weeks)	-0.5487	-0.1031	-0.1130	0.0225	
	(0.1545)	(0.0367)	(0.0581)	(0.0510)	
Family head married (=1 if yes w/ spouse present)	-0.0529	0.0975	-0.0212	0.0717	
	(0.0345)	(0.0279)	(0.0428)	(0.0363)	
Covered by health insurance last month (=1 if yes)	-0.0546	-0.0432	-0.0012	0.0655	
	(0.0302)	(0.0285)	(0.0421)	(0.0381)	
Live in Midwest region (=1 if yes)	0.0813	0.0909	0.0694	0.0696	
	(0.0176)	(0.0181)	(0.0243)	(0.0243)	
Live in South region (=1 if yes)	0.0371	0.0535	0.2847	0.3558	
	(0.0606)	(0.0617)	(0.1995)	(0.2031)	
ive in West region (=1 if yes)	-0.0025	-0.0022	-0.0111	-0.0134	
	(0.0035)	(0.0036)	(0.0071)	(0.0073)	
Live in Metropolitan Area (=1 if yes)	-0.0941	-0.0748	-0.1876	-0.2018	
	(0.0242)	(0.0245)	(0.0312)	(0.0318)	
Child's sex	-0.0519	-0.0436	-0.0519	-0.0552	
	(0.0279)	(0.0286)	(0.0351)	(0.0363)	
'hild's age	0.0001	0.0002			
-	(0.0014)	(0.0015)			
'hild's age squared	-3.14E-08	-1.04E-07			
	(1.43E-06)	(1.49E-06)			
Black (=1 if yes)	-0.0526	-0.0542			
	(0.0222)	(0.0228)			
Hispanic (=1 if yes)	-0.0074	-0.0001	-0.0490	-0.0388	
	(0.0336)	(0.0340)	(0.0450)	(0.0460)	
10ther's age at child's birth	-0.0147	-0.0168	0.0283	0.0307	
	(0.0311)	(0.0316)	(0.0415)	(0.0428)	
Iother's age at child's birth sq	0.0170	0.0183	0.0333	0.0219	
······································	(0.0347)	(0.0355)	(0.0458)	(0.0475)	
Mother smoked while pregnant (=1 if yes)	-0.0124	-0.0033	0.0003	-0.0020	
	(0.0190)	(0.0195)	(0.0259)	(0.0265)	
Constant	6.1247	2.3544	0.5419	-0.8406	
	(1.3555)	(0.3625)	(1.4461)	(1.4293)	
Sample size	2,463	2,320	1,597	1,482	
R-squared	0.1078	0.1096	0.1835	0.1983	

Notes: standard errors are in parenthesis, regional omitted category: northeast

(#) This variable is the predicted value from te first step estimation presented in table 7. In models (9) and (11) the relevant IV is child's height and in models (10) and (12) the relevant IV is mother's BMI.

Table 9 Descriptive statistic, selected variables, clasiffied according to "family food insecurity"

	Children	Children between 6 and 11 years old			Children between 12 and 16 years old		
Variable	Family with food insecurity	Family without food insecurity	Pooled	Family with food insecurity	Family without food insecurity	Pooled	
Inputs for possible instruments							
School with fixed price full breakfast (=1 if yes)	0.77***	0.65***	0.66	0.78***	0.63***	0.65	
	(0.02)	(0.01)	(0.01)	(0.03)	(0.01)	(0.01)	
Times per week child eats breakfast at school	3.55***	2.33***	2.49	2.62***	1.62***	1.75	
	(0.12)	(0.05)	(0.05)	(0.18)	(0.06)	(0.06)	
School with fixed price full lunch (=1 if yes)	1.00	0.96	0.96	1.01**	0.97**	0.97	
	(0.02)	(0.01)	(0.01)	(0.04)	(0.01)	(0.01)	
Times per week child eats breakfast at school	4.71***	4.06***	4.13	4.28***	3.74***	3.80	
	(0.06)	(0.03)	(0.03)	(0.10)	(0.05)	(0.04)	
Number of observations (##)	497	4,207	4,704	242	1,967	2,209	

Note: Standard errors are in parenthesis. * indicates a statistically significant difference between food-insecure and food-secure families at a 90% confidence interval, ** at a 95% and *** at a 99% confidence interval
(##) The total number of observations for each variable may vary due to missing cases
Source: NHANES III Youth and Exam data files